A Markov chain model for daily rainfall occurrences at east Thanjavur district

M. THIYAGARAJAN, RAMADOSS and RAMARAJ

A.V.V.M. Sri Pushpam College, Poondi, Tamil Nadu (Received 7 April 1994, Modified 2 June 1995)

खार — वर्षा होने और न होने के बारे में दिस्थितीय एक मार्कोव चेन की सहायता से बताया जा सकता है। वर्षा होने की स्थिति को 0 स्थिति डारा तथा वर्षा न होने की स्थिति को 1 स्थिति डारा दर्शाया गया है। ज्ञात प्राचल के साथ प्वासों प्रक्रिया का अनुसरण करने वाला नमूना हमने श्वपनाया है। हस प्वासों नमूने का उपयोग करते हुए मार्कोव चेन के सिदांतों और भूतकालिक अप्रोक्षित मानों अथवा अब तक अनप्रोक्षत भावी मानों से संबंधित स्थिति का आकलन करने डेतु साँख्यिकी निष्कर्षों को प्रभावित करने के लिए डमने एक नया तरीका अपनाया है। इस शोध का उद्देश्य नए दृष्टिकोण के साथ आंकड़ों की पूर्व उपयुक्तता का तुलनात्मक अध्ययन करना है।

ABST_{bACT}. The occurrences and non-occurrences of the rainfall can be described by a two-state Markov chain. A dry date is denoted by state 0 and wet date is denoted by state 1. We have taken the sample which follows a Poisson process with known parameter. Using this Poisson sample we have given a new approach to affect statistical inference for the law of the Markov chain and state estimation concerning unobserved past values or not yet observed future values. The paper aims at comparing the earlier fit of the data with the new approach.

Key words - Markov chain, Probability, Dry and wet days, Matrix, Monsoon.

1. Introduction

Karr (1986) has given the following procedure to problems on inference for Markov chains based on point process samples :

Given a process $X = [X(t)]$ on R with unknown probability law, but usually of fairly narrowly specified structure, and a Poisson process N on R , often but not invariably preassumed homogeneous, such that X and N are independent, the goal is to affect statistical inference for the law of X and state estimation concerning unobserved past values or not yet observed (and possibly never to be observed) future values, given as of single realizations of the Poisson samples $X(T_n)$, that is, the values of X at the points of N , together possibly with observations of N itself. Observable aspects are in the marked point process N , marked by the current value of X .

We follow the assumptions made by Karr (1986) in applying Markov chain model for studying the patterns of occurrences of dry and wet days. The days were recorded under two categories, dry and wet. A day is called dry if the rainfall during the 24 hours commencing from 6 a.m. on that day is either nil or trace, otherwise it is termed wet. One of the two periods is pre-monsoon from 1 January to 30 June and the other period is monsoon season from 1

July to 30 December. Two sets of pre-monsoon season are considered: Set I covering a total $2 \times 180 = 360$ days in the two years - 1987 and 1988. The data are taken from records maintained by the revenue officials of Government of Tamil Nadu at the concerned centres.

In redrafting the information in a table suitable to Karr's (1986) discussion, data are chosen to be a Poisson process with parameter 2 and to decide the observation to fall on specific dates in every month.

2. Method and analysis

To describe the occurrences of dry and wet days at Thanjavur district we use a two-state Markov chain model as considered by Gabriel and Neumann (1983). We assume that the probability of dry (or wet) days depends on the conditions of previous day. Such a probability model is referred to as a Markov chain, whose parameters are the two conditional probabilities,

 P_0 = Prob (Wet day/previous day dry)

 P_1 = Prob (Wet day/previous day wet)

We denote the dry day by state 0 and wet day as state 1, the occurrence of dry and wet day can, then,

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TABLE 1

Centre: Thiruthuraipoondi - Years: 1987-1988

TABLE 2

Centre: Vedaranyam - Years: 1987-1988

	Jan		Feb		Mar		Apr		May		Jun		Jul		Aug		Sep		Oct		Nov		Dec	
$Day -$	1987 1988 1987 1988 1987 1988 1987 1988 1987 1988 1987 1988 1987 1988 1987 1988 1987 1988 1987 1988 1987 1988																							
4	D	D	D	D	D	D	D	D	W	D	D	D	D	D	D	D	D	D	W	D	D	W	W	W
5	D	D	D	D	D	D	D	D	D	D	D	D	D	D	D	D	D	D	D	D	W	W	W	D
10	D	D	D	D	D	D	D	D	D	D	D	D	D	W	D	D	D	D	W	W	W	W	D	W
11	D	W	D	D	D	D	D	D	D	D	D	D	D	D	W	W	D	D	W	D	D	W	W	D
18	D	D	D	D	D	D	D	W	D	D	D	D	D	D	W	D	D	D	W	D	W	D	D	D
25	D	D	D	D	D	D	D	W	D	W	D	D	D	D	D	W	D	D	W	D	D	D	W	D
29	D	D		D	D	D	D	D	D	D	D	D	D	D	\overline{D}	W	W	D	w	D	D	D	D	D
30	D	Ð				D	D	D	D	D	Ð	D		D	\Box	W	W	D	W	D	D	D	D	W

 $\frac{\partial}{\partial t}$

 \mathbf{P}

be denoted by a Markov chain with two ergodic states 0 and 1 and transition probability matrix given by.

$$
A = \begin{pmatrix} 0 & 1 \\ 1 - p_0 & p_0 \\ 1 & 1 - p_1 \end{pmatrix}
$$

The *n*-step transition probabilities are given by the elements of $Aⁿ$, where $Aⁿ$ is given by,

$$
A^{n} = \frac{1}{p_0 + p_1} \begin{pmatrix} p_1 & p_0 \\ p_1 & p_0 \end{pmatrix} + \frac{(1 - p_0 - p_1)^n}{p_0 + p_1} \begin{pmatrix} p_0 & -p_0 \\ -p_1 & p_1 \end{pmatrix}
$$
\n\nFurther
$$
\lim_{n \to \infty} A^{n} = \begin{pmatrix} v_0 & v_1 \\ v_0 & v_1 \end{pmatrix}
$$

where
$$
v_0 = \frac{p_1}{p_0 + p_1}
$$
, $v_1 = \frac{p_0}{p_0 + p_1}$

A wet spell W of length k is defined as a sequence of k wet days preceded and followed by $\frac{dy}{dx}$ days. Dry spells *D* are defined correspondingly. Weather cycles are defined as combinations of a wet spell and an adjacent dry spell. It follows that the distribution of W and D are geometric with parameters p_1 and p_0 respectively, that is Probability $(W=r) = p_1 (1-p_1)^{r-1}$, $r = 1, 2, ...$ and
Probability $(D=r) = p_0 (1-p_0)^{r-1}$ so that the mean
 $E(W) = \frac{1}{p_1}$ and $E(D) = \frac{1}{p_0}$

The central limit theorem for dependent variables asserts that the distribution of the number of wet days, Y_n in a sequence of *n* transitions is asymptotically normal with,

$$
E(Y_n) = \frac{np_0}{p_0 + p_1}, \text{ VAR } (Y_n) = np_0 p_1 \frac{(2 - p_0 - p_1)}{(p_0 + p_1)^3}
$$

From Table 1, we observe that

$$
A = \begin{pmatrix} 0.91176 & 0.08824 \\ 0.5 & 0.5 \end{pmatrix}
$$

$$
A^{20} = \begin{pmatrix} 0.85003 & 0.15002 \\ 0.85002 & 0.15002 \end{pmatrix}
$$

$$
n \rightarrow \infty
$$

$$
A^{n} = (0.85, 0.15)
$$

Actual day	Total		
Total	day	Total	
0	1		
Total	8	7	15
Total	70	13	83
$A = \begin{pmatrix} 0.91176 & 0.08824 \\ 0.53333 & 0.46667 \end{pmatrix}$			
$A^{20} = \begin{pmatrix} 0.85804 & 0.14196 \\ 0.85804 & 0.14196 \\ 0.85804 & 0.14196 \end{pmatrix}$			

$$
n \to \infty
$$
 $A^n = (0.858, 0.142)$

From Table 2, we observe that

Actual day
\nTotal
\n0 1
\n1 8 9 17
\nTotal
\nTotal
\n66 16 82
\n
$$
A = \begin{pmatrix} 0.89231 & 0.10769 \\ 0.47059 & 0.52941 \end{pmatrix}
$$
\n
$$
A^{20} = \begin{pmatrix} 0.81577 & 0.18712 \\ 0.82030 & 0.18816 \end{pmatrix}
$$
\n
$$
n \rightarrow \infty
$$
\n
$$
A^{n} = (0.816, 0.187)
$$

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TABLE 3

Centre: Nagapattinam - Years: 1987-1988

TABLE 4

Limiting distribution

ä,

TABLE 5

Mean and variance of wet spell

Centres		Thiruthuraipoondi		Vedaranyam	Nagapattinam		
Years	1987	1988	1987	1988	1987	1988	
$E(Y_n)$	4.5	4.26	5.59	5.42	7.02	6.33	
$VAR(Y_n)$	9.18	8.1	11.18	7.69	13.85	9.98	

From Table 3, we observe that

Lt $n\rightarrow\infty$ $A^n=(0.766,\ 0.234)$

3. Conclusion

We find that the Markov Chain of order one gives an adequate fit for the observations of the premonsoon and the monsoon period during 1987-88. This model is based on the sample which follows a Poisson process from the observations obtained throughout the year. The limiting distribution of the transit from wet to dry and dry to wet is shown in Table 4.

The distribution of the number of wet days Y_n in a sequence of n transitions is asymptotically normal with mean and variance as shown in Table 5. We find that Markov chain model, fitted from out of the Poisson samples of the data, gives a better picture of the situation than that obtained by Thiyagarajan (1989) based on raw data. Statistical

test shows that the Markov chain of order 1 gives an adequate fit for the observations of the premonsoon period (1987-88).

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